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AN A-CONTRARIO APPROACH TO QUASI-PERIODIC NOISE REMOVAL

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ABSTRACT

Images can be affected by quasi-periodic noise. This undesirable feature manifests itself by spurious repetitive patterns covering the whole image, well localized in the Fourier domain. While notch filtering permits to get rid of this phenomenon, this however requires to first detect the resulting Fourier spikes, and, in particular, to discriminate between noise spikes and spectrum patterns caused by spatially localized textures or repetitive structures. This paper proposes a statistical a-contrario detection of noise spikes in the Fourier domain. A Matlab code is also provided.

Index Terms— quasi-periodic noise, a-contrario method.

1. INTRODUCTION

Images may be affected by quasi-periodic noise, that is, spurious repetitive patterns covering the entire image. This artifact is often caused by electrical interferences during image acquisition or transmission, which make remote sensing applications especially prone to the phenomenon. Periodic noise gives more or less sharp spikes in the image spectrum, which can be filtered out using notch filters. Some basic approaches requiring expert tuning can be found in [1, 2]. The difficulty is to automate spike detection, that is, notch filter design. Some authors [3, 4] suggest to detect spikes in the Fourier domain as large deviations with respect to a localized median value. However, distinguishing between spikes caused by a localized texture or a repetitive structure (common in man-made environments) and spurious ones caused by periodic noise is still challenging. It has been observed in [5] that periodic noise is likely to be the only periodic structure present in any patch extracted from the impaired image. It is proposed to design the notch filter based on the average power spectrum computed on a set of patches: the contribution of localized repetitive textures is smoothed out by averaging, and the only remaining spikes are the ones caused by periodic noise. We have recently automated this approach by detecting the spikes in the averaged spectrum as statistical outliers of the distribution expected from natural (non-noisy) patches, which is known to follow the inverse of a power of the frequency [6].

We propose to detect the Fourier spikes caused by periodic noise with an a-contrario method. A-contrario the-

ory was introduced for the detection of alignments in images [7] and proves to be well adapted to many image analysis tasks as Gestalt grouping [8, 9], detection of moving objects in videos [10], sub-pixel change detection in satellite imagery [11], shape identification [12], line segment [13] or elliptical arc [14] detection, interest point matching [15, 16, 17, 18], point clustering [19], multiple object detection [20], or spot detection in textures [21], just to cite a few recent papers. This detection theory is based on the idea that features of interest (called *meaningful* features) are not likely to be caused by a random background process. Deciding whether a feature is meaningful or not is based on the *number of false alarms* (NFA) which corresponds to the average number of such a feature expected from the background process (hence “false alarms”). More precisely, in the case of real-valued features, if the features of interest are not likely to have a high value x , and provided that the meaningful features are sought among N features, then the NFA of the observed x is:

$$\text{NFA}(x) = N \Pr(X \geq x) \quad (1)$$

where $\Pr(X \geq x)$ is the probability that a random feature X following the background process is larger than x . In a-contrario models, X is actually a function of independent variables (although recent papers explore non-independence [22, 23]), $\Pr(X > x)$ is thus calculated from marginal laws, which are either parametric [8] or empirically estimated [12]. Once an NFA has been defined, the feature of interest are in most papers those such that $\text{NFA} \leq 1$. This means that at most one such feature could be observed in the background model among the N tested ones. In this sense, a-contrario detection is a parameter-free method. Psychophysical experiments suggest that a-contrario detection matches human perception [24, 25].

2. A-CONTRARIO DETECTION OF SPECTRUM SPIKES CAUSED BY PERIODIC NOISE

We adopt the same basic premise as in [5, 6]: considering a set of P patches of size L spanning the entire image, periodic noise is characterized by spikes present in the spectrum of every patch. In contrast, an image not affected by periodic noise does not show this feature, and localized periodic textures give Fourier spikes only in some of the patches. Our aim is thus to detect frequencies whose corresponding Fourier coefficients have an unexpected large amplitude in *all*

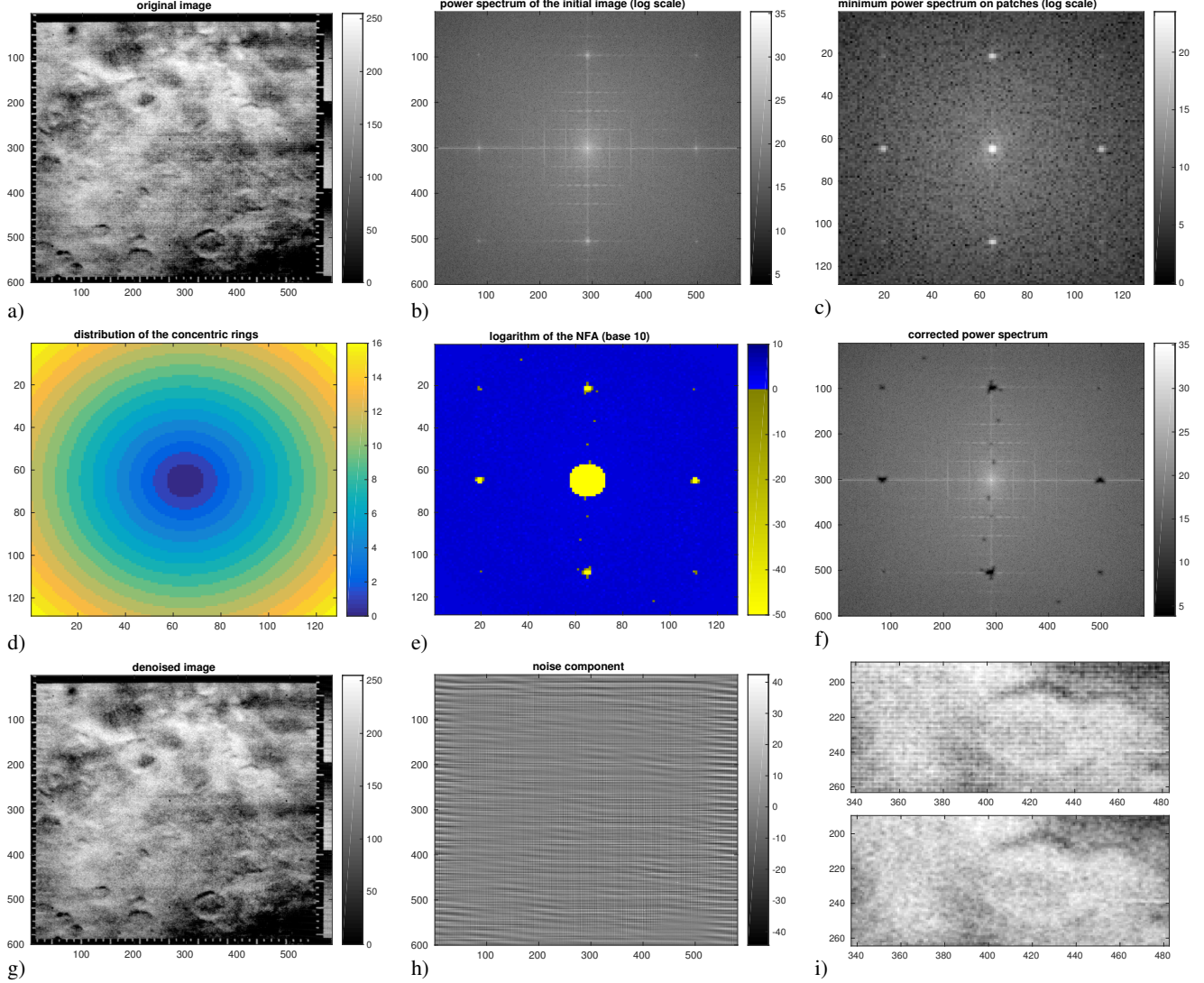


Fig. 1. a) Original image; b) spectrum; c) element-wise minimum of the spectra of the 128×128 patches; d) distribution of the 16 concentric rings; e) map of the NFA; f) original power spectrum with the periodic noise spikes smoothed out; g) denoised image; h) retrieved noise component; i) comparison between close-ups of the original and denoised images.

patches, with an a-contrario approach. Let us consider the Fourier coefficient $c_{n,m}^p$ of frequency (n, m) cycles per image of the p -th patch. We would like to characterize (n, m) for which the minimum value $\min_{1 \leq p \leq P} |c_{n,m}^p|$ is not likely to be observed in a non-noisy image (here $|\cdot|$ is the complex modulus). In this case, there is a suspicious spike at (n, m) in *all* of the P patches. Let us define the background process for any (n, m) as $X_{n,m} = \min_{1 \leq p \leq P} |C_{n,m}^p|$ with the $C_{n,m}^p$ independent and identically distributed random variables, their tail distribution being noted $F_{n,m}$. Because of the independence assumption, the probability in (1) writes

$$\Pr \left(\min_{1 \leq p \leq P} |C_{n,m}^p| \geq x \right) = \Pr (\forall p, |C_{n,m}^p| \geq x) \quad (2)$$

$$= (F_{n,m}(x))^P \quad (3)$$

The problem is now to design the distribution of the $C_{n,m}^p$ in the background model, i.e., for a non-noisy image. This distribution cannot obviously be chosen identical for all (n, m) but should instead at least depend on the frequency $f = \sqrt{n^2 + m^2}$, the most striking evidence being the expected $1/f$ power law of the spectrum [26]. We thus choose to cover the spectrum (centered at low frequencies) by $R = L/8$ concentric rings, as in Fig. 1d), the distribution being fixed over a given ring but varying from ring to ring. Each distribution is empirically estimated from the Fourier coefficients of the P patches belonging to the considered ring. That is, for any x ,

$$F_{\mathcal{R}(n,m)}(x) = \frac{1}{P \# \mathcal{R}(n,m)} \# \left\{ (n', m', p) \right. \quad (4)$$

$$\left. \text{s.t. } |c_{n',m'}^p| \geq x, (n', m') \in \mathcal{R}(n, m), 1 \leq p \leq P \right\}$$

where $\#\cdot$ denotes the cardinality of any finite set, and $\mathcal{R}(n, m)$ denotes the ring to which (n, m) belongs.

Note that it is easy to add an angular discretization to Fig. 1d) in order to satisfy anisotropic models as in [27].

In a given ring, only a small minority of coefficients may belong to a spike caused by periodic noise. It is thus valid to consider that $F_{\mathcal{R}(n, m)}(x)$ is the probability that $|C_{n, m}| \geq x$ for a non-noisy image, in accordance with the background process. The expected number of false alarms over a given ring \mathcal{R} (where the tail distribution is $F_{\mathcal{R}}$) is thus $\#\mathcal{R} \cdot F_{\mathcal{R}}(x)$. Since there are R such rings,

$$\text{NFA}(|c_{n, m}|) = R \#\mathcal{R}(n, m) (F_{\mathcal{R}(n, m)}(|c_{n, m}|))^P \quad (5)$$

is a valid definition for a number of false alarms (see Def. 4 and Prop. 2 in [21]), in the sense that it satisfies:

Proposition *The expected number of random spectrum coefficients following the background process with an NFA below ε is smaller than ε , i.e.,*

$$E(\#\{(n, m) \text{ s.t. } \text{NFA}(|C_{n, m}|) \leq \varepsilon\}) \leq \varepsilon \quad (6)$$

Consequently, frequencies (n, m) such that $\min_p c_{n, m}^p$ has an NFA below 1 are not likely to come from a non-noisy image, and are confidently considered as coming from periodic noise.

Remark 1 A-contrario detection of suspicious spectrum structures is also used in [28] for image aliasing and in [29] for motion blur, without, however, any spectrum model.

Remark 2 In [5, 6], noise spikes are detected from an average power spectrum. With the notation of (2)-(3), the distribution of $X_{n, m} = \frac{1}{P} \sum_{p=1}^P |C_{n, m}^p|$ is a scaled P -fold convolution of the distribution of $|C_{n, m}|$ with itself. An NFA could thus be defined similarly to (2)-(3), in the same spirit as in [18]. It could benefit from the central limit theorem by approximating the distribution of $X_{n, m}$ for large P by a normal distribution, of mean value $\simeq 1/f^\gamma$. Similarly, the empirical $F_{n, m}^P$ in (3) could be replaced with a parametric distribution such as a Gumbel, Fréchet or Weibull distribution, following the extreme value theory [30].

3. QUASI-PERIODIC NOISE REMOVAL

Spikes caused by periodic noise are detected by thresholding the map of the $\text{NFA}(|c_{n, m}|)$ to 1. Based on these spikes, the same strategy as in [6] is employed: the spikes are propagated to the original image spectrum and permits to build a notch filter which smooths out the contribution of the periodic noise. Given an image impaired by quasi-periodic noise, the proposed algorithm is described in Alg. 1. We use the same implementation details as in [6] where a thorough discussion of the parameters can also be found.

Fig. 1 shows a running example. a) is an image obtained by the Mariner 4 probe, affected by a quasi-periodic noise, and b) is its power spectrum. Visually discriminating the noise components in the spectrum is not obvious. c) shows the minimum of the power spectra from the patches (step 2, here $L = 128$ pixels). Four spikes are clearly visible, in

Algorithm 1 A-contrario quasi-periodic noise removal

Input: image i (size $X \times Y$, Fourier transform I), patch size L .

1. Extract non-overlapping patches (to enforce the independence assumption in (2)-(3)) of size $L \times L$ distributed over i (giving P patches) and calculate the $F_{\mathcal{R}}$ with (4).
 2. Calculate the minimum of the power spectra of the patches: for any (n, m) , $|c_{n, m}| = \min_{p=1 \dots P} |c_{n, m}^p|$.
 3. For any (n, m) , calculate $\text{NFA}(|c_{n, m}|)$ with (5).
 4. Define the spike map M_o^P on the $L \times L$ spectrum such that $M_o^P(n, m) = 1$ if $\text{NFA}(|c_{n, m}|) \leq 1$, and 0 otherwise.
 5. Interpolate the outlier map M_o^P of size $L \times L$ to $X \times Y$, giving a map M_o of the probable spurious spikes in the original image spectrum. Multiplying the initial image spectrum by $1 - M_o$ acts as a notch filter, eliminating the influence of the quasi-periodic noise.
 6. Retrieve \hat{n} , estimation of the periodic noise component, as the inverse Fourier transform (IFT) of $M_o I$, and \hat{i} , estimation of the denoised image, as $i - \hat{n}$ (i.e., IFT of $(1 - M_o)I$).
-

addition to the central low-frequency component. d) shows the distribution of the $R = L/8 = 16$ concentric rings on which the Fourier coefficient amplitudes are assumed to be identically distributed. The central disk (in deep blue) is not taken into account by the algorithm, because low frequencies do not correspond to repetitive patterns (moreover, the identical distribution hypothesis is not valid in this disk). e) is the map of the logarithm of the NFA (with an arbitrary value in the central disk): a large majority of the coefficients give an NFA larger than 1 ($\log(\text{NFA}) > 0$), except for the four spikes with a very low NFA ($\simeq 10^{-180}$). f) is the corrected spectrum $(1 - M_o)I$ (step 5): we can see that, in addition to the four spikes, a few other coefficients have been detected as meaningful (with $\text{NFA} \simeq 10^{-1}$). g) is \hat{i} , and h) is \hat{n} (step 6). Real image details cannot be seen in \hat{n} , which suggests that the noise component has been well separated from the actual data. i) shows close-ups of i and \hat{i} .

4. EXPERIMENTS

Some results are presented in Fig. 2. First, a synthetic periodic noise has been added to the Mandrill image. The underlying image and the noise component are well separated by our algorithm. This shows that the proposed a-contrario detection is able to detect the noise spikes even in the presence of a strong high-frequency texture (the fur here) affecting most patches. The two remaining experiments deal with real images. The Lunar Orbiter image is impaired by striping, which is adequately removed. The Mariner 6 image is affected by noise patterns whose spectrum is much more complex than just a few separated spikes. The algorithm is still able to separate the noise component from the image.

Additional experiments and a Matlab code are available at: www.loria.fr/%7Eesur/software/ACARPENOS/

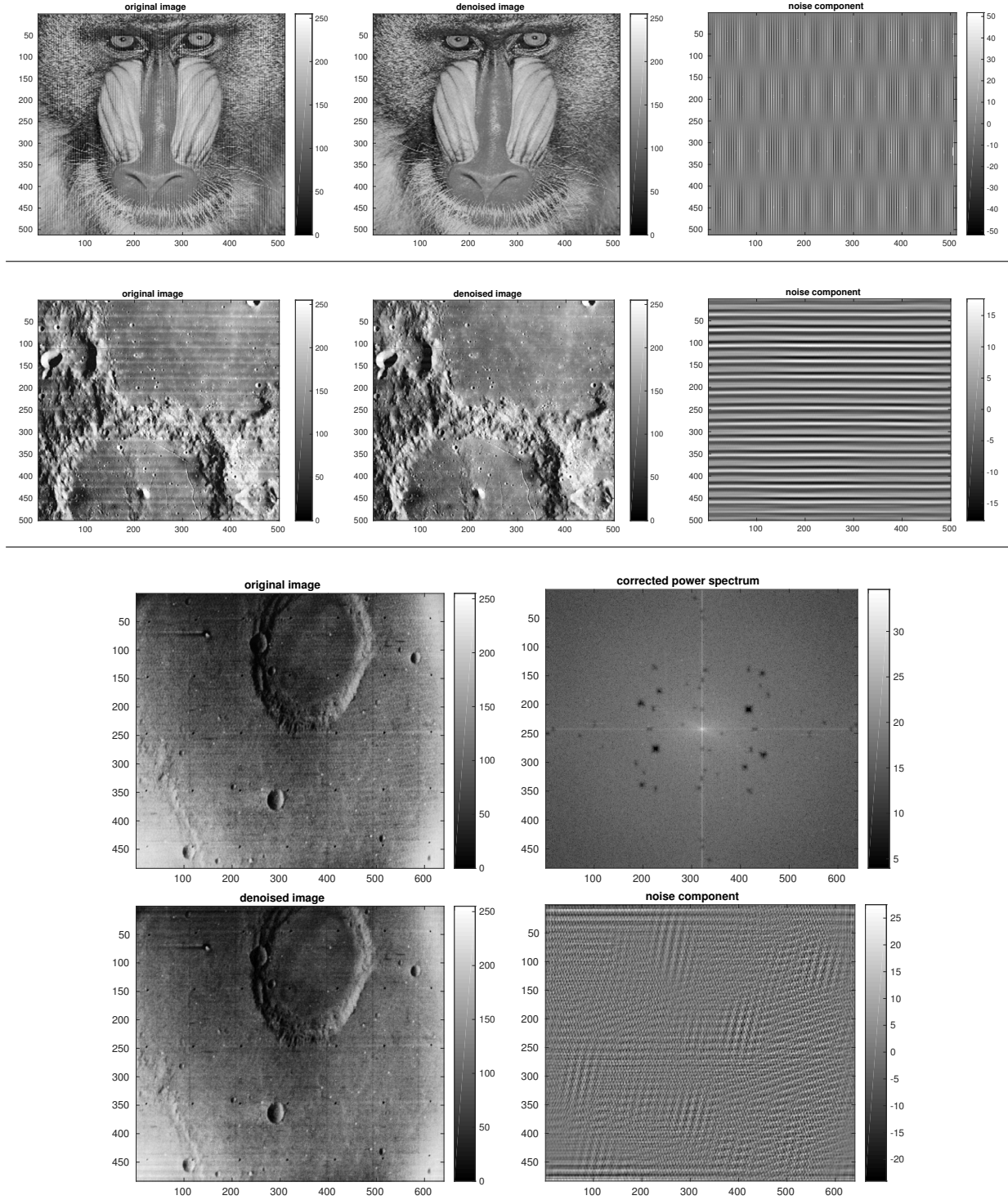


Fig. 2. From top to bottom: Mandrill, Lunar Orbiter, and Mariner 6 experiments. The reader is asked to zoom in on the pdf file.

5. CONCLUSION

This paper discussed a new a-contrario detection of the spectrum components of a quasi-periodic noise, making it possible to design notch filters for periodic noise removal. While ex-

isting similar methods [5, 6] use an average power spectrum calculated from image patches, the proposed method is based on the minimum spectrum and benefits from the parameter-free decision method offered by the a-contrario approach.

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